



IT Solutions for  
Animal Production

# Computing strategies for a single step SNP model with an across country reference population

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# OUTLINE

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# Introduction: current genomic evaluation models

- Genomic model (Meuwissen et al. 2001) revolutionises animal breeding, particularly for Holsteins
- A multiple step genomic model for German Holsteins
  - Conventional bull proofs deregressed
    - National and international MACE evaluations
  - Deregressed bull proofs for SNP effect estimation with an across-country genomic reference population (EuroGenomics)
  - Male pedigree index excluded overestimated EBV of bull dams
  - DGV of candidates combined with conventional male pedigree index
- Advantages and drawbacks of the current multi-step genomic model
  - Simple for implementation
  - Genomic reference population customised as wished
    - Only progeny-tested bulls with a minimum EDC
    - No cows included due to possible overestimated EBV
  - Conventional EBV will be biased by genomic pre-selection



## Introduction: single step genomic model

- Invention of  $\mathbf{H}^{-1}$  matrix (Misztal et al., Christensen & Lund)
  - Accurate integration of genotyped animals into conventional evaluation

$$\mathbf{H} = \begin{bmatrix} \mathbf{A}_{11} & \mathbf{A}_{12} \\ \mathbf{A}_{21} & \mathbf{G}_{22} \end{bmatrix} = \mathbf{A} + \begin{bmatrix} \mathbf{0} & \mathbf{0} \\ \mathbf{0} & \mathbf{G}_{22} - \mathbf{A}_{22} \end{bmatrix}$$

$$\mathbf{H}^{-1} = \begin{bmatrix} \mathbf{A}^{11} & \mathbf{A}^{12} \\ \mathbf{A}^{21} & \mathbf{G}_{22}^{-1} + \mathbf{A}^{22} - \mathbf{A}_{22}^{-1} \end{bmatrix} = \mathbf{A}^{-1} + \begin{bmatrix} \mathbf{0} & \mathbf{0} \\ \mathbf{0} & \mathbf{G}_{22}^{-1} - \mathbf{A}_{22}^{-1} \end{bmatrix}$$

- Computing strategies for large populations (Legarra & Ducrocq)
- Single step GBLUP model works perfectly for closed populations with all genotypes and phenotypes in one hand (Misztal et al.)
- Further developments for 'open' Holstein genomic evaluations
  - Using deregressed MACE EBV of foreign reference bulls
    - as substitutes of original phenotype data of foreign cows
  - Efficient interim genomic evaluations without new phenotypes
    - Instead of running the whole genomic evaluation

Reducing the impact of inflated EBV of genotyped cows



## A single step SNP model (Goddard & Liu, 2012)

- A mixed linear model in a general form

$$\mathbf{y} = \mathbf{X}\mathbf{b} + \mathbf{Z}_p\mathbf{p} + \mathbf{W}\mathbf{u} + \mathbf{e} \quad \text{var}(\mathbf{p}) = \mathbf{I}\sigma_p^2 \quad \text{var}(\mathbf{e}) = \mathbf{I}\sigma_e^2$$

- For genotyped animals (group 2)

$$\mathbf{u}_2 = \mathbf{Z}\mathbf{g} + \mathbf{a}_2$$

- Distribution of SNP marker effects (BLUP or Bayesian models)

- $\text{var}(\mathbf{g}) = \mathbf{B}\sigma_g^2$  (e.g. BLUP SNP model:  $\mathbf{B} = b\mathbf{I} = \frac{1-k}{m}\mathbf{I}$ )

- Residual polygenic effects of genotyped animals

$$\text{var}(\mathbf{a}_2) = \mathbf{A}_{22}k\sigma_g^2 \quad \text{var}(\mathbf{u}_2) = \mathbf{G}_{22}\sigma_g^2 = (\mathbf{Z}\mathbf{B}\mathbf{Z}' + k\mathbf{A}_{22})\sigma_g^2$$

- Conditional distribution for non-genotyped animals (group 1)

$$\mathbf{u}_1 = \mathbf{T}\mathbf{u}_2 + \mathbf{d} \quad \text{with transmission matrix } \mathbf{T} = \mathbf{A}_{12}\mathbf{A}_{22}^{-1}$$

$$\text{a deviation effect } \text{Var}(\mathbf{d}) = \mathbf{D}\sigma_g^2$$

- Joint distribution for genotyped and non-genotyped animals

$$\mathbf{G} = \text{Var}(\mathbf{u}) = \text{Var} \begin{pmatrix} \mathbf{u}_1 \\ \mathbf{u}_2 \end{pmatrix} = \begin{bmatrix} \mathbf{T}\mathbf{G}_{22}\mathbf{T}' + \mathbf{D} & \mathbf{T}\mathbf{G}_{22} \\ \mathbf{G}_{22}\mathbf{T}' & \mathbf{G}_{22} \end{bmatrix} \sigma_g^2$$



## SSS model (Goddard & Liu, 2012)

- Inverse of (co)variance matrix for  $\mathbf{u}$

$$\mathbf{G}^{-1} = \begin{bmatrix} \mathbf{D}^{-1} & -\mathbf{D}^{-1}\mathbf{T} \\ -\mathbf{T}'\mathbf{D}^{-1} & \mathbf{G}_{22}^{-1} + \mathbf{T}'\mathbf{D}\mathbf{T} \end{bmatrix} \sigma_g^{-2} \quad \mathbf{G}^{-1} = \begin{bmatrix} \mathbf{A}^{11} & \mathbf{A}^{12} \\ \mathbf{A}^{21} & \mathbf{G}_{22}^{-1} + \mathbf{A}^{22} - \mathbf{A}_{22}^{-1} \end{bmatrix} \sigma_g^{-2}$$

- Joint distribution of  $\mathbf{u}$  and SNP effects

$$\mathbf{H} = \text{var} \begin{bmatrix} \mathbf{u}_1 \\ \mathbf{u}_2 \\ \mathbf{g} \end{bmatrix} = \begin{bmatrix} \mathbf{T}\mathbf{G}_{22}\mathbf{T}' + \mathbf{D} & \mathbf{T}\mathbf{G}_{22} & \mathbf{T}\mathbf{Z}\mathbf{B} \\ \mathbf{G}_{22}\mathbf{T}' & \mathbf{G}_{22} & \mathbf{Z}\mathbf{B} \\ \mathbf{B}\mathbf{Z}'\mathbf{T}' & \mathbf{B}\mathbf{Z}' & \mathbf{B} \end{bmatrix} \sigma_g^2$$

$$\mathbf{H}^{-1} = \begin{bmatrix} \mathbf{D}^{-1} & -\mathbf{D}^{-1}\mathbf{T} & \mathbf{0} \\ -\mathbf{T}'\mathbf{D}^{-1} & \frac{1}{k}\mathbf{A}_{22}^{-1} + \mathbf{T}'\mathbf{D}^{-1}\mathbf{T} & -\frac{1}{k}\mathbf{A}_{22}^{-1}\mathbf{Z} \\ \mathbf{0} & -\frac{1}{k}\mathbf{Z}'\mathbf{A}_{22}^{-1} & \mathbf{B}^{-1} + \frac{1}{k}\mathbf{Z}'\mathbf{A}_{22}^{-1}\mathbf{Z} \end{bmatrix} \sigma_g^{-2}$$

$$= \begin{bmatrix} \mathbf{A}^{11} & \mathbf{A}^{12} & \mathbf{0} \\ \mathbf{A}^{21} & \mathbf{A}^{22} + (\frac{1}{k} - 1)\mathbf{A}_{22}^{-1} & -\frac{1}{k}\mathbf{A}_{22}^{-1}\mathbf{Z} \\ \mathbf{0} & -\frac{1}{k}\mathbf{Z}'\mathbf{A}_{22}^{-1} & \mathbf{B}^{-1} + \frac{1}{k}\mathbf{Z}'\mathbf{A}_{22}^{-1}\mathbf{Z} \end{bmatrix} \sigma_g^{-2}$$



## SSS model (Goddard & Liu, 2012)

- Mixed model equations for all the effects

$$\begin{bmatrix} \mathbf{X}'\mathbf{X} & \mathbf{X}'\mathbf{Z}_p & \mathbf{X}'\mathbf{W} \\ \mathbf{Z}_p'\mathbf{X} & \mathbf{Z}_p'\mathbf{Z}_p + \mathbf{I}\delta & \mathbf{Z}_p'\mathbf{W} \\ \mathbf{W}'\mathbf{X} & \mathbf{W}'\mathbf{Z}_p & \mathbf{W}'\mathbf{W} + \mathbf{H}^{-1}\sigma_e^2 \end{bmatrix} \begin{bmatrix} \hat{\mathbf{b}} \\ \hat{\mathbf{p}} \\ \hat{\mathbf{u}} \end{bmatrix} = \begin{bmatrix} \mathbf{X}'\mathbf{y} \\ \mathbf{Z}_p'\mathbf{y} \\ \mathbf{W}'\mathbf{y} \end{bmatrix}$$

$$\begin{bmatrix} \mathbf{X}'\mathbf{X} & \mathbf{X}'\mathbf{Z}_p & \mathbf{X}'\mathbf{W}_1 & \mathbf{X}'\mathbf{W}_2 & \mathbf{0} & \mathbf{0} \\ \mathbf{Z}_p'\mathbf{X} & \mathbf{Z}_p'\mathbf{Z}_p + \mathbf{I}\delta & \mathbf{Z}_p'\mathbf{W}_1 & \mathbf{Z}_p'\mathbf{W}_2 & \mathbf{0} & \mathbf{0} \\ \mathbf{W}_1'\mathbf{X} & \mathbf{W}_1'\mathbf{Z}_p & \mathbf{W}_1'\mathbf{W}_1 + \lambda\mathbf{A}^{11} & \lambda\mathbf{A}^{12} & \mathbf{0} & \mathbf{0} \\ \mathbf{W}_2'\mathbf{X} & \mathbf{W}_2'\mathbf{Z}_p & \lambda\mathbf{A}^{21} & \mathbf{W}_2'\mathbf{W}_2 + \lambda(\mathbf{A}^{22} + (\frac{1}{k}-1)\mathbf{A}_{22}^{-1}) & -\frac{1}{k}\lambda\mathbf{A}_{22}^{-1}\mathbf{Z} & \mathbf{0} \\ \mathbf{0} & \mathbf{0} & \mathbf{0} & -\frac{1}{k}\lambda\mathbf{Z}'\mathbf{A}_{22}^{-1} & \lambda(\mathbf{B}^{-1} + \frac{1}{k}\mathbf{Z}'\mathbf{A}_{22}^{-1}\mathbf{Z}) & \mathbf{0} \end{bmatrix} \begin{bmatrix} \hat{\mathbf{b}} \\ \hat{\mathbf{p}} \\ \hat{\mathbf{u}}_1 \\ \hat{\mathbf{u}}_2 \\ \hat{\mathbf{g}} \end{bmatrix} = \begin{bmatrix} \mathbf{X}'\mathbf{y} \\ \mathbf{Z}_p'\mathbf{y} \\ \mathbf{W}_1'\mathbf{y} \\ \mathbf{W}_2'\mathbf{y} \\ \mathbf{0} \end{bmatrix}$$

- Solve two sets of equations iteratively:

$$\begin{bmatrix} \mathbf{X}'\mathbf{X} & \mathbf{X}'\mathbf{Z}_p & \mathbf{X}'\mathbf{W}_1 & \mathbf{X}'\mathbf{W}_2 \\ \mathbf{Z}_p'\mathbf{X} & \mathbf{Z}_p'\mathbf{Z}_p + \mathbf{I}\delta & \mathbf{Z}_p'\mathbf{W}_1 & \mathbf{Z}_p'\mathbf{W}_2 \\ \mathbf{W}_1'\mathbf{X} & \mathbf{W}_1'\mathbf{Z}_p & \mathbf{W}_1'\mathbf{W}_1 + \lambda\mathbf{A}^{11} & \lambda\mathbf{A}^{12} \\ \mathbf{W}_2'\mathbf{X} & \mathbf{W}_2'\mathbf{Z}_p & \lambda\mathbf{A}^{21} & \mathbf{W}_2'\mathbf{W}_2 + \lambda(\mathbf{A}^{22} + (\frac{1}{k}-1)\mathbf{A}_{22}^{-1}) \end{bmatrix} \begin{bmatrix} \hat{\mathbf{b}} \\ \hat{\mathbf{p}} \\ \hat{\mathbf{u}}_1 \\ \hat{\mathbf{u}}_2 \end{bmatrix} = \begin{bmatrix} \mathbf{X}'\mathbf{y} \\ \mathbf{Z}_p'\mathbf{y} \\ \mathbf{W}_1'\mathbf{y} \\ \mathbf{W}_2'\mathbf{y} + \frac{1}{k}\lambda\mathbf{A}_{22}^{-1}\mathbf{Z}\hat{\mathbf{g}} \end{bmatrix}$$

$$(\mathbf{B}^{-1} + \frac{1}{k}\mathbf{Z}'\mathbf{A}_{22}^{-1}\mathbf{Z})\hat{\mathbf{g}} = \frac{1}{k}\mathbf{Z}'\mathbf{A}_{22}^{-1}\hat{\mathbf{u}}_2$$



## Computing strategies for the SSS model

- Re-arranging equation for  $\mathbf{u}_2$

$$\mathbf{W}_2' \mathbf{X} \hat{\mathbf{b}} + \mathbf{W}_2' \mathbf{Z}_p \hat{\mathbf{p}} + \lambda \mathbf{A}^{21} \hat{\mathbf{u}}_1 + (\mathbf{W}_2' \mathbf{W}_2 + \lambda \mathbf{A}^{22}) \hat{\mathbf{u}}_2 = \mathbf{W}_2' \mathbf{y} + \lambda \mathbf{A}_{22}^{-1} (\hat{\mathbf{u}}_2 - \frac{1}{k} \hat{\mathbf{a}}_2)$$

- New part on top of conventional MME: pure genomic contribution

$$\mathbf{A}_{22}^{-1} (\hat{\mathbf{u}}_2 - \frac{1}{k} \hat{\mathbf{a}}_2) = \mathbf{A}_{22}^{-1} \hat{\mathbf{u}}_2^* = \hat{\boldsymbol{\varphi}}$$

- Re-arranging SNP equations

$$\mathbf{B}^{-1} \hat{\mathbf{g}} = \frac{1}{k} \mathbf{Z}' \mathbf{A}_{22}^{-1} (\hat{\mathbf{u}}_2 - \mathbf{Z} \hat{\mathbf{g}}) = \frac{1}{k} \mathbf{Z}' \mathbf{A}_{22}^{-1} \hat{\mathbf{a}}_2 = \frac{1}{k} \mathbf{Z}' \hat{\boldsymbol{\gamma}}$$

$$\hat{\mathbf{g}} = \frac{1}{k} \mathbf{B} \mathbf{Z}' \hat{\boldsymbol{\gamma}}$$

- Two core calculations  $\hat{\boldsymbol{\varphi}} = \mathbf{A}_{22}^{-1} \hat{\mathbf{u}}_2^*$  and  $\hat{\boldsymbol{\gamma}} = \mathbf{A}_{22}^{-1} \hat{\mathbf{a}}_2$  can be done by solving equations (Legarra & Ducrocq, 2012) with Gauss-Jacobi (VanRaden)

$$\mathbf{A}_{22} \hat{\boldsymbol{\varphi}} = \hat{\mathbf{u}}_2^* \quad \mathbf{A}_{22} \hat{\boldsymbol{\gamma}} = \hat{\mathbf{a}}_2$$

- A direct algorithm for computing  $\boldsymbol{\varphi}$  and  $\boldsymbol{\gamma}$  (Liu & Goddard)

- Calculating 'special' EBV of non-genotyped relatives
- Additional decomposing  $\mathbf{A}^{-1*}$  (besides  $\mathbf{A}^*$ ) using Colleau's method
- No setup of  $\mathbf{A}_{22}$  or  $\mathbf{A}_{22}^{-1}$  needed





## Features of the SSS model

- A simple and closed form of  $\mathbf{H}^{-1}$  (including SNP effects)
- No large matrix or product of large matrices in MME
- No need for genomic relationship matrix  $\mathbf{G}$  or  $\mathbf{G}^{-1}$  or  $\mathbf{A}_{22}$  or  $\mathbf{A}_{22}^{-1}$ 
  - No limits on genotyped animals
- Suited for *Iteration on Data* technique for populations of any size
- Flexible SNP effect modelling: Bayesian or BLUP SNP models
  - One step Bayesian model (Goddard & Liu 2012)
- A residual polygenic effect in the SSS model
  - Analogue to SSGBlup using weighted  $\mathbf{G}$  matrix:  $\text{var}(\mathbf{u}_2) = \mathbf{G}_{22} \sigma_g^2 = (\mathbf{Z}\mathbf{B}\mathbf{Z}' + k\mathbf{A}_{22})\sigma_g^2$
  - Removed overestimation bias of genomic prediction (Liu et al. 2011)
  - Numerical equivalence:  $k=0.0001$  as no residual polygenic effect
  - $k=0.9999$  as no SNP/DGV effects
  - RPG connects genotyped animals to phenotyped population
  - Removed large matrix multiplications (e.g.  $\mathbf{Z}'\mathbf{A}^{22}\mathbf{Z}$ ,  $\mathbf{Z}'\mathbf{A}^{21}$ )
  - Residual polygenic variance estimated or determined via genomic validation
- Similar to Gengler's model (EAAP 2012) but with a different derivation



## SNP effect estimation: reference population

- Single step GBLUP has no SNP effect estimation step and thus no direct control of information flow from reference pop. to candidates
- But genomic prediction can be improved, if RP is controlled by:
  - Removing bulls with limited data & less reliable EBV (with biases)
  - Deleting bull dams or cows with preferential treatments
  - Genotyped candidates without phenotypes or with imputed genotypes

Animals of other breeds in a multi-breed evaluation

- Introduce a **filter**:  $\mathbf{F} = \text{diag}\{1, 0, 0, 1, \dots, 1, 0\}$  to SNP equation:

$$\hat{\mathbf{g}} = \frac{1}{k} \mathbf{BZ}' \mathbf{A}_{22}^{-1} \hat{\mathbf{a}}_2 \longrightarrow \hat{\mathbf{g}} = \frac{1}{k} \mathbf{BZ}' \mathbf{F} \mathbf{A}_{22}^{-1} \hat{\mathbf{a}}_2$$

- 1 / 0: genotyped animal is included / excluded in reference population
  - 0.9 for imputed genotypes from a low density chip
- Meanwhile keep ALL animals (ref. or not) in  $\mathbf{u}$ ,  $\mathbf{a}_2$  and  $\mathbf{A}_{22}$

Impact of genomic pre-selection on  $\mathbf{u}$  is not influenced by the selection of reference animals



# Estimate SNP effects with a special algorithm

A 'large  $p$  and small  $n$ ' computational problem

An efficient Gauss-Seidel algorithm with a special residual update (GSRU, Legarra & Misztal 2008)

- For a given set of  $\mathbf{u}_2$  estimates:

$$\mathbf{Z}(\hat{\mathbf{g}}^{[j+1]} - \hat{\mathbf{g}}^{[j]}) = (\hat{\mathbf{u}}_2 - \hat{\mathbf{a}}_2^{[j+1]}) - (\hat{\mathbf{u}}_2 - \hat{\mathbf{a}}_2^{[j]}) = \hat{\mathbf{a}}_2^{[j]} - \hat{\mathbf{a}}_2^{[j+1]}$$

$$\hat{\mathbf{a}}_2^{[j+1]} = \hat{\mathbf{a}}_2^{[j]} - \mathbf{Z}(\hat{\mathbf{g}}^{[j+1]} - \hat{\mathbf{g}}^{[j]})$$

- An efficient estimation procedure:

- At an outer iteration round calculate for ALL genotyped animals

$$\hat{\mathbf{a}}_2 = \hat{\mathbf{u}}_2 - \mathbf{Z}\hat{\mathbf{g}}$$

- An inner loop ( $j$ -th round) for separating SNP from RPG effects

- Step 1. estimate SNP effects

$$\hat{\mathbf{g}}^{[j+1]} = \frac{1}{k} \mathbf{BZ}' \mathbf{F} \mathbf{A}_{22}^{-1} \hat{\mathbf{a}}_2^{[j]}$$

- Step 2. update residual polygenic effects

$$\hat{\mathbf{a}}_2^{[j+1]} = \hat{\mathbf{a}}_2^{[j]} - \mathbf{Z}(\hat{\mathbf{g}}^{[j+1]} - \hat{\mathbf{g}}^{[j]})$$



# Interim genomic evaluation w/o new phenotypes

- In contrast to conventional evaluation, genomic prediction is a more continuous process, monthly, weekly, or on-demand (just-in-time)
- Using SNP effect estimates can easily provide genomic evaluations as genotypes available any time between two major evaluations
- **Simple formulae for GEBV** instead of running the complete system
- Equation  $\mathbf{u}_2$  is simplified ( $\mathbf{y} = \mathbf{0}$  for candidates)

$$\mathbf{W}_2' \mathbf{X} \hat{\mathbf{b}} + \mathbf{W}_2' \mathbf{Z}_p \hat{\mathbf{p}} + \lambda \mathbf{A}^{21} \hat{\mathbf{u}}_1 + (\mathbf{W}_2' \mathbf{W}_2 + \lambda (\mathbf{A}^{22} + (\frac{1}{k} - 1) \mathbf{A}_{22}^{-1})) \hat{\mathbf{u}}_2 = \mathbf{W}_2' \mathbf{y} + \frac{1}{k} \lambda \mathbf{A}_{22}^{-1} \mathbf{Z} \hat{\mathbf{g}}$$



$$\mathbf{A}^{21} \hat{\mathbf{u}}_1 + \mathbf{A}^{22} \hat{\mathbf{u}}_2 - \mathbf{A}_{22}^{-1} (\hat{\mathbf{u}}_2 - \frac{1}{k} \hat{\mathbf{a}}_2) = \mathbf{0}$$

- A candidate ( $l$ ) has sire ( $s$ ) and dam ( $d$ )

$$u_l = \frac{A}{A+B} \mathbf{z}' \mathbf{g} + \frac{d_l}{A+B} \frac{1}{2} (u_s + u_d) + \frac{1}{A+B} \sum_{m=1}^n a_{22}^{lm} (u_m - \frac{1}{k} a_m)$$

SNP/DGV effects

Parental average

Correction for genotyped relatives



## Interim genomic evaluation w/o new phenotypes

- A general formula for GEBV of candidate  $l$

$$u_l = \frac{A}{A+B} \mathbf{z}' \mathbf{g} + \frac{d_l}{A+B} \frac{1}{2} (u_s + u_d) + \frac{1}{A+B} \sum_{m=1}^n a_{22}^{lm} (u_m - \frac{1}{k} a_m)$$

- If all ancestors (both parents) of the candidate are genotyped (appr.)

$$u_l = \mathbf{z}' \mathbf{g} + \frac{d_l^*}{1+d_l^*} \frac{1}{2} (a_s + a_d)$$

- Genotyped relatives influence the correction term:

$$\frac{1}{A+B} \sum_{m=1}^n a_{22}^{lm} (u_m - \frac{1}{k} a_m)$$

- A good approximation using nucleus family: genotyped sire, dam/MGS, (direct progeny and mate) of the genotyped animal
- Alternative: select index method may be used to combine DGV and parental average as in case of multiple step genomic model

Assumption: contribution of new candidate genotypes to GEBV of genotyped population is negligible



# Genomic and phenotypic data for DEU Holsteins



- Pedigree data for national and international evaluations
  - 76 million animals in **vit** database
  - 571,000 animals in Interbull Holstein bull pedigree
- Phenotypic test-day data (milk yield, August 2013 evaluation)
  - 19 mln DEU cows with test-day records (25 mln animals in pedigree)
  - 340 mln test-day records
- MACE phenotype (August 2013 MACE evaluation)
  - 133,028 Holstein bulls (representing 70 mln cows worldwide)
- Genotype data (45,613 SNPs selected from Illumina 50K v2)
  - 93,233 genotyped animals (278,000 animals in pedigree)
    - 6978 cows with test-day data
    - 26,361 Holstein bulls in EuroGenomics RP (c.a. 34 mln cows)
      - 18,497 bulls with only or more MACE info
- Combined (inter)national genotype, phenotype and pedigree data sets for genomic evaluation using the SSS model



## Integration of MACE with test-day data

- A three-lactation random regression test-day model for each of three production traits for German Holsteins
- DEU random regression model uses Legendre polynomials with three terms:  $u_i = t_1 * c_{1i} + t_2 * c_{2i} + t_3 * c_{3i}$
- Every animal has 3 lactations x 3 coefficients = 9 EBVs
- But: MACE phenotype is a single trait deregressed proof (DRP) on a combined lactation basis ( $EBV_{comb} = w_1 * u_1 + w_2 * u_2 + w_3 * u_3$ )
- Majority of reference bulls (> 67% in case of DEU Holstein) have only one single DRP available for genomic evaluation
  - Little info for SNP effect in form of random regression coefficients
  - **SNP effect on a combined lactation basis**
- However, all animals, including candidates and foreign bulls without domestic daughters, have GEBV expressed in random regression coefficient form



# SSS & test-day models: trait definition changes

- SNP effect estimation with a single trait model
  - Majority of reference bulls have a single MACE EBV
  - **Condense** 3 x 3 = 9 RRC per animal to 1 combined lactation EBV

$$\hat{\mathbf{g}} = \frac{1}{k} \mathbf{BZ}' \mathbf{F} \hat{\boldsymbol{\gamma}} = \frac{1}{k} \mathbf{BZ}' \mathbf{F} \mathbf{A}_{22}^{-1} \hat{\mathbf{a}}_2$$

- Adjusting conventional MME for genomic contribution

- Conventional MME is expressed on a 3 x 3 = 9 RRC basis  $\hat{\mathbf{u}}_2 = \sum_{l=1}^3 \sum_{m=1}^3 t_m c_{ij}$
- **Expand**  $\mathbf{z}'\mathbf{g}$  (DGV) from one combined lactation to RRC basis

$\delta = \mathbf{z}'\mathbf{g} / \hat{\mathbf{u}}_2$  on a single trait (combined lactation) basis

RRC for  $\hat{\mathbf{u}}_2 : c_{ij}$

RRC for  $DGV(\mathbf{z}'\mathbf{g}) = \delta c_{ij}$

Genotypes do not change lactation curve shapes, only curve areas

Lactation curve shapes determined only by phenotypes





## Conclusions and Summary

- Our single step SNP model provides
  - Useful SNP effects estimates
  - Flexible SNP effect modelling (Bayes or BLUP, diagonal matrix **B**)
  - No need for **G** or **G**<sup>-1</sup>, no direct setup for **A**<sub>22</sub> or **A**<sub>22</sub><sup>-1</sup>
  - Suited for iteration on data techniques because of no large matrices
- A residual polygenic effect for reducing prediction bias
- Unlike SSGblup, our SSS model has a SNP effect estimation step
  - Identical modelling GEBV as SSGblup
- Avoids bias in evaluations caused by genomic pre-selection
- Active control of genomic information flow from reference population to candidates by incl./excl. animals from reference population
- Simple formulae for frequent interim genomic evaluations
- Applicable for 'open' Holstein system with mixed (inter)national phenotypes, genotypes and pedigree

Reliability approximation using genomic relationships



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- Mike Goddard for deriving the single step SNP model
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**THANK YOU!**



**vit**



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## SSS model without RPG (Goddard & Liu, 2012)

- For genotyped animals  $\hat{u}_2 = Z\hat{g}$
- Mixed model equations for all the effects

$$\begin{bmatrix} X'X & X'Z_p & X'W_1 & 0 \\ Z_p'X & Z_p'Z_p + I\delta & Z_p'W_1 & 0 \\ W_1'X & W_1'Z_p & W_1'W_1 + \lambda A^{11} & \lambda A^{12}Z \\ 0 & 0 & \lambda Z'A^{21} & Z'W_2'W_2Z + \lambda Z'(A^{22} - A_{22}^{-1})Z + \lambda B^{-1} \end{bmatrix} \begin{bmatrix} \hat{b} \\ \hat{p} \\ \hat{u}_1 \\ \hat{g} \end{bmatrix} = \begin{bmatrix} X'y \\ Z_p'y \\ W_1'y \\ Z'W_2'y \end{bmatrix}$$

- Solve two sets of equations iteratively:

$$\begin{bmatrix} X'X & X'Z_p & X'W_1 \\ Z_p'X & Z_p'Z_p + I\delta & Z_p'W_1 \\ W_1'X & W_1'Z_p & W_1'W_1 + \lambda A^{11} \end{bmatrix} \begin{bmatrix} \hat{b} \\ \hat{p} \\ \hat{u}_1 \end{bmatrix} = \begin{bmatrix} X'y \\ Z_p'y \\ W_1'y - \lambda A^{12}Z\hat{g} \end{bmatrix}$$

$$(Z'W_2'W_2Z + \lambda Z'(A^{22} - A_{22}^{-1})Z + \lambda B^{-1})\hat{g} = Z'W_2'y - \lambda Z'A^{21}\hat{u}_1$$

- Similar to equations (Legarra & Ducrocq, 2012)



# Comparing single step SNP & SSGblup models

- Legarra and Ducrocq (2012)

- Genomic contribution to RHS of MME

$$\alpha_u \hat{\phi} - \alpha_u \hat{\gamma} = \alpha_u \mathbf{A}_{22}^{-1} \hat{\mathbf{u}}_2 - \alpha_u \mathbf{G}^{-1} \hat{\mathbf{u}}_2 \quad [1]$$

- SNP effect estimation (Strandén & Garrick 2009)

$$\hat{\mathbf{g}} = \mathbf{DZ}' \mathbf{G}^{-1} \hat{\mathbf{u}}_2 \quad [2]$$

- Another formulation using the terms by Goddard & Liu (2012)

$$\lambda \hat{\phi} - \lambda \hat{\gamma} = \lambda \mathbf{A}_{22}^{-1} \hat{\mathbf{u}}_2 - \lambda \mathbf{G}_{22}^{-1} \hat{\mathbf{u}}_2 \quad [3]$$

$$\hat{\mathbf{g}} = \mathbf{BZ}' \mathbf{G}_{22}^{-1} \hat{\mathbf{u}}_2 \quad [4]$$

- Corresponding equations by Goddard and Liu (2012)

- Genomic contribution to RHS of MME

$$\lambda \mathbf{A}_{22}^{-1} \hat{\mathbf{u}}_2 - \frac{1}{k} \lambda \mathbf{A}_{22}^{-1} \hat{\mathbf{a}}_2 \quad [5]$$

- SNP effect estimation

$$\hat{\mathbf{g}} = \frac{1}{k} \mathbf{BZ}' \mathbf{A}_{22}^{-1} \hat{\mathbf{a}}_2 \quad [6]$$



## Algorithm for calculating $\varphi_2 = \mathbf{A}_{22}^{-1} \mathbf{u}_2$

- Misztal et al. (2009) and Ducrocq et al. (2012) transformed the matrix inversion into an equation solving issue:  $\mathbf{A}_{22} \varphi_2 = \mathbf{u}_2$

- Adding relatives of genotyped animals to the equation  $\mathbf{u}_2$  :

$$\begin{bmatrix} \mathbf{A}_{00} & \mathbf{A}_{02} \\ \mathbf{A}_{20} & \mathbf{A}_{22} \end{bmatrix} \begin{bmatrix} \mathbf{0} \\ \varphi_2 \end{bmatrix} = \begin{bmatrix} \mathbf{u}_0 \\ \mathbf{u}_2 \end{bmatrix} \longrightarrow \mathbf{A} \hat{\varphi} = \mathbf{u}$$

- where  $\begin{bmatrix} \mathbf{A}_{00} & \mathbf{A}_{02} \end{bmatrix} \begin{bmatrix} \mathbf{0} \\ \varphi_2 \end{bmatrix} = \mathbf{u}_0 \longrightarrow \mathbf{u}_0 = \mathbf{A}_{02} \varphi_2$

- Inverting the complete relationship matrix  $\mathbf{A}$  (Mrode, 2005):

$$\hat{\varphi} = \mathbf{T}^{-T} (\mathbf{D}^{-1} \mathbf{T}^{-1} \mathbf{u})$$

- Solve the equations by reading pedigree twice (Colleau, 2002)

$$\hat{\eta} = \mathbf{D}^{-1} \mathbf{T}^{-1} \mathbf{u}$$

$$\hat{\eta}_i = d^i [u_i - \frac{1}{2}(u_s + u_d)]$$

from oldest to youngest

$$\hat{\varphi} = \mathbf{T}^{-T} \hat{\eta}$$

$$\hat{\varphi}_i = \hat{\varphi}_i + \hat{\eta}_i$$

$$\hat{\varphi}_s = \hat{\varphi}_s + (-\frac{1}{2})\hat{\eta}_i$$

$$\hat{\varphi}_d = \hat{\varphi}_d + (-\frac{1}{2})\hat{\eta}_i$$

from youngest to oldest



## Algorithm for $\mathbf{u}_0$ for non-genotyped relatives

- Calculating  $\mathbf{u}_0$  for non-genotyped relatives

$$\mathbf{u} = \begin{bmatrix} \mathbf{u}_0 \\ \mathbf{u}_2 \end{bmatrix} = \mathbf{A}\hat{\boldsymbol{\phi}} = \mathbf{TDT}'\hat{\boldsymbol{\phi}} = \mathbf{TDT}' \begin{bmatrix} 0 \\ \hat{\boldsymbol{\phi}}_2 \end{bmatrix}$$

- Solve equations by reading pedigree twice (Colleau, 2002)

$$\hat{\boldsymbol{\zeta}} = \mathbf{T}'\hat{\boldsymbol{\phi}}$$

$$\mathbf{u} = \mathbf{TD}\hat{\boldsymbol{\zeta}}$$

$$\hat{\xi}_i = \hat{\xi}_i + \hat{\phi}_i$$

$$\hat{\xi}_s = \hat{\xi}_s + \frac{1}{2}\hat{\xi}_i$$

$$\hat{\xi}_d = \hat{\xi}_d + \frac{1}{2}\hat{\xi}_i$$

from youngest to oldest

$$u_i = d_i \hat{\xi}_i + \frac{1}{2}(u_s + u_d)$$

from oldest to youngest



## Algorithms for calculating $\varphi_2 = \mathbf{A}_{22}^{-1} \mathbf{u}_2$

- An iterative solving procedure (Misztal et al, Legarra & Ducrocq)

$$\mathbf{A} \hat{\boldsymbol{\varphi}}^* = \mathbf{u} \qquad \mathbf{v} = \mathbf{T} \mathbf{D} \mathbf{T}' \tilde{\boldsymbol{\varphi}}^*$$

- Gauss-Jacobi solving:  $\varphi_j = ((\mathbf{u}_2 - \mathbf{v}_2)_j + \text{diag}(\mathbf{A}_{22})_j * \tilde{\varphi}_j) / \text{diag}(\mathbf{A}_{22})_j$   
(VanRaden, personal communication, 2012)
- Straightforward, not using  $\mathbf{A}^{-1}$ .

- A direct solving procedure using the inverse relationship matrix

- 1. choose starting values for genotyped animals, eg.  $\tilde{\boldsymbol{\varphi}}_2 = \mathbf{u}_2$
- 2. calculate for non-genotyped relatives:  $\mathbf{u}_0 = \mathbf{A}_{02} \boldsymbol{\varphi}_2$
- 3. estimate using  $\mathbf{A}^{-1}$ :  $\hat{\boldsymbol{\varphi}} = \mathbf{T}^{-T} (\mathbf{D}^{-1} \mathbf{T}^{-1} \mathbf{u})$
- 4.  $\boldsymbol{\varphi}_2^{[i]} = w \boldsymbol{\varphi}_2^{[i]} + (1 - w) \boldsymbol{\varphi}_2^{[i-1]}$  e.g.  $w=0.5$
- 5. check convergence of  $\hat{\boldsymbol{\varphi}}_2$  only for genotyped animals
- 6. if not converged, repeat steps 2 to 4 until converged

- The direct procedure may be optimised by estimating  $w$ .





## Reliability approximation for SSS model

- Single step GBLUP model provides direct info than SSS model
  - Properly scaled genomic relationship matrix  $\mathbf{G}_{22}$
- Genomic added value:  $\varpi [\mathbf{G}_{22}^{-1} - \mathbf{A}_{22}^{-1}]$ 
  - Obtained from genomic validations
- Approximation method avoiding matrix inversion and even direct forming both matrices in core
- One unified reliability procedure for ALL groups of animals
- Theoretical genomic reliability is adjusted to realised genomic reliability
- Currently, most countries assume ONE single constant genomic EDC for pure genomic contribution
- Candidates with sires having no phenotype should have lower reliability than those with sires in reference population
- Reliability approximation also for interim genomic evaluation



## Adjusting overestimated cow EBV

- Current multi-step genomic model uses DRP of bulls
- Overestimated bull dams have little impact on SNP effects
- Male pedigree index of candidates do not use bull dam EBV directly
- In single step genomic model, phenotypes of bull dams can no longer be excluded
  - Using filter **F** to exclude genotyped bull dams from SNP effect estimation
  - Impact of overestimated bull dam EBV still exists in candidate GEBV
- Inflated bull dams seem to exist only in production traits
- Strategy for adjusting possible inflated bull dam EBV
  - Identify cows or bull dams as potentially preferentially treated
  - Fit special fixed lactation curves to test-day data



## Model differences between national and MACE

- MACE evaluation uses a single trait model on bull level
- DEU national evaluations apply multi-trait models on animal level
  - a three-lactation random regression test-day model (MFP+SCS)
  - a multiple trait fertility model
  - a three-parity animal model with direct and maternal effects (calving)
  - a multiple trait model for conformation and workability traits
  - a non-linear survival model for direct functional longevity
- Integrating (previous) MACE evaluation of ALL (foreign) bulls into single step genomic model is technically challenging
- Foreign bulls enter the SSS system as animals with records
- Modify least squares part of LHS and RHS of mixed model equations to integrate MACE phenotypic information
- Bulls in 3 groups with only national phenotypic information, only foreign data, and both



## Three groups of bulls with different data sources



- Conventional phenotypic information per bull
  - EDC:  $\varphi^{\text{NAT}}$ ,  $\varphi^{\text{MACE}}$ ,
  - Deregressed MACE proof: DRP
- Add one equation to mixed model equations per bull
- ONLY foreign daughter phenotypes are to be added to MME
- Group 1: bulls without daughters/progeny in home country
  - LHS =  $\varphi^{\text{MACE}}$ , RHS =  $\varphi^{\text{MACE}} * \text{DRP}$  ( $\varphi^{\text{NAT}}=0$ )
- Group 2: bulls with daughters/progeny ONLY in home country: add nothing ( $\varphi^{\text{MACE}} - \varphi^{\text{NAT}} = 0$ )
- Group 3: bulls with daughters/progeny in BOTH home and foreign countries
  - LHS =  $\varphi^{\text{MACE}} - \varphi^{\text{NAT}}$ , RHS =  $(\varphi^{\text{MACE}} - \varphi^{\text{NAT}}) * \text{DRP}$



## Discussion: genomic pre-selection

- Evaluation bias due to genomic pre-selection can be avoided only when all culled candidates are also considered (Patry & Ducrocq)
- Equation  $\mathbf{u}_2$  is simplified ( $\mathbf{y} = \mathbf{0}$ ) for candidates

$$\mathbf{W}_2' \mathbf{X} \hat{\mathbf{b}} + \mathbf{W}_2' \mathbf{Z}_p \hat{\mathbf{p}} + \lambda \mathbf{A}^{21} \hat{\mathbf{u}}_1 + (\mathbf{W}_2' \mathbf{W}_2 + \lambda (\mathbf{A}^{22} + (\frac{1}{k} - 1) \mathbf{A}_{22}^{-1})) \hat{\mathbf{u}}_2 = \mathbf{W}_2' \mathbf{y} + \frac{1}{k} \lambda \mathbf{A}_{22}^{-1} \mathbf{Z} \hat{\mathbf{g}}$$

$$\mathbf{A}^{21} \hat{\mathbf{u}}_1 + \mathbf{A}^{22} \hat{\mathbf{u}}_2 - \mathbf{A}_{22}^{-1} \hat{\mathbf{u}}_2 + \frac{1}{k} \mathbf{A}_{22}^{-1} \hat{\mathbf{u}}_2 - \frac{1}{k} \mathbf{A}_{22}^{-1} \mathbf{Z} \hat{\mathbf{g}} = \mathbf{0}$$

$$\mathbf{A}^{21} \hat{\mathbf{u}}_1 + \mathbf{A}^{22} \hat{\mathbf{u}}_2 = \mathbf{A}_{22}^{-1} (\hat{\mathbf{u}}_2 - \frac{1}{k} \hat{\mathbf{a}}_2)$$

$$\mathbf{A}^{21} \hat{\mathbf{u}}_1 + \mathbf{A}^{22} \hat{\mathbf{u}}_2 = \mathbf{A}_{22}^{-1} \hat{\mathbf{u}}_2^* = \hat{\phi}$$

- Pure genomic contribution  $\hat{\phi}$  of the candidates is propagated to all relatives via pedigree
- In fact, this process for candidates applies also to genotyped sires of domestic cows with raw records
- But the candidates can be excluded from SNP effect estimation

# Interim genomic evaluation w/o new phenotypes



- Monthly, weekly or on-demand (just-in-time) genomic evaluations between two major evaluations (no new phenotypes available)
- Equation  $\mathbf{u}_2$  is simplified ( $\mathbf{y} = \mathbf{0}$ )

$$\mathbf{W}_2' \mathbf{X} \hat{\mathbf{b}} + \mathbf{W}_2' \mathbf{Z}_p \hat{\mathbf{p}} + \lambda \mathbf{A}^{21} \hat{\mathbf{u}}_1 + (\mathbf{W}_2' \mathbf{W}_2 + \lambda (\mathbf{A}^{22} + (\frac{1}{k} - 1) \mathbf{A}_{22}^{-1})) \hat{\mathbf{u}}_2 = \mathbf{W}_2' \mathbf{y} + \frac{1}{k} \lambda \mathbf{A}_{22}^{-1} \mathbf{Z} \hat{\mathbf{g}}$$

$$\mathbf{A}^{21} \hat{\mathbf{u}}_1 + \mathbf{A}^{22} \hat{\mathbf{u}}_2 - \mathbf{A}_{22}^{-1} \hat{\mathbf{u}}_2 + \frac{1}{k} \mathbf{A}_{22}^{-1} \hat{\mathbf{u}}_2 - \frac{1}{k} \mathbf{A}_{22}^{-1} \mathbf{Z} \hat{\mathbf{g}} = \mathbf{0}$$

$$\mathbf{A}^{21} \hat{\mathbf{u}}_1 + \mathbf{A}^{22} \hat{\mathbf{u}}_2 - \mathbf{A}_{22}^{-1} (\hat{\mathbf{u}}_2 - \frac{1}{k} \hat{\mathbf{a}}_2) = \mathbf{0}$$

- A candidate ( $l$ ) has sire ( $s$ ) and dam ( $d$ )

$$\text{diag}(\mathbf{A}^{22})_i = 1 + d_i \quad 1/d_i = 0.5 - (f_s + f_d)/4$$

$$\text{diag}(\mathbf{A}_{22}^{-1})_i = 1 + d_i^*$$

$$[(d_i - d_i^*) + \frac{1}{k}(1 + d_i^*)] u_i - \frac{1}{2} d_i (u_s + u_d) - \frac{1}{k} (1 + d_i^*) \mathbf{z}' \mathbf{g} + \sum_{m=1}^n a_{22}^{lm} (-u_m + \frac{1}{k} a_m) = 0$$

$$[(d_i - d_i^*) + \frac{1}{k}(1 + d_i^*)] u_i = \frac{1}{k} (1 + d_i^*) \mathbf{z}' \mathbf{g} + \frac{1}{2} d_i (u_s + u_d) + \sum_{m=1}^n a_{22}^{lm} (u_m - \frac{1}{k} a_m) = 0$$

$$u_i = \frac{A}{A+B} \mathbf{z}' \mathbf{g} + \frac{d_i}{A+B} \frac{1}{2} (u_s + u_d) + \frac{1}{A+B} \sum_{m=1}^n a_{22}^{lm} (u_m - \frac{1}{k} a_m)$$

**SNP/DGV effects**      **Parental average**      **Correction for genotyped relatives**

$$A = \frac{1}{k} (1 + d_i^*) \quad B = d_i - d_i^* \quad \text{(small variance)}$$



## SSS & test-day models: solving strategy

- Foreign bulls without domestic daughters
  - Entering the SSS model as animals with own data
  - Only one single MACE EBV/DRP on a combined lactation basis
  - **Starting values** for random regression coefficients
    - $q_{\text{comb}} = q_1 = q_2 = q_3$
    - $q_i = t_1 * c_1 + t_2 * c_2 + t_3 * c_3 = t_1 * c_1$
    - $c_{1i} = q_{\text{comb}} / t_1$  and  $c_{2i} = c_{3i} = 0$
- Procedures for solving MME of the SSS model
  - Iterate conventional MME for some rounds or using solution priors
  - Add MACE phenotypes of foreign bulls
    - Single trait model on combined lactation basis:  $\text{DRP} = \mu + \text{EBV} + \varepsilon$
  - Add genomic correction term of RHS
 
$$\mathbf{W}_2' \mathbf{X} \hat{\mathbf{b}} + \mathbf{W}_2' \mathbf{Z}_p \hat{\mathbf{p}} + \lambda \mathbf{A}^{21} \hat{\mathbf{u}}_1 + (\mathbf{W}_2' \mathbf{W}_2 + \lambda \mathbf{A}^{22}) \hat{\mathbf{u}}_2 = \mathbf{W}_2' \mathbf{y} + \lambda \mathbf{A}_{22}^{-1} (\hat{\mathbf{u}}_2 - \frac{1}{k} \hat{\mathbf{a}}_2)$$
  - Convert  $\mathbf{u}_2$  to SNP effects



## Estimate SNP effects with a special algorithm

- A 'large  $p$  and small  $n$ ' computational problem
- An efficient Gauss-Seidel algorithm with a special residual update (GSRU, Legarra & Misztal 2008)

- For a given set of  $\mathbf{u}_2$  estimates:

$$\mathbf{Z}(\hat{\mathbf{g}}_i^{[j+1]} - \hat{\mathbf{g}}_i^{[j]}) = (\hat{\mathbf{u}}_2 - \hat{\mathbf{a}}_{2(+i)}^{[j]}) - (\hat{\mathbf{u}}_2 - \hat{\mathbf{a}}_{2(-i)}^{[j]}) = \hat{\mathbf{a}}_{2(-i)}^{[j]} - \hat{\mathbf{a}}_{2(+i)}^{[j]}$$

- Computing procedure:

- At  $j$ -th iteration round calculate for ALL genotyped animals

$$\hat{\mathbf{a}}_2^{[j]} = \hat{\mathbf{u}}_2^{[j]} - \mathbf{Z}\hat{\mathbf{g}}^{[j]}$$

- An inner loop over SNP ( $i = 1, \dots, m$ ) sorted by heterozygosity (D)

- Step 1. 
$$\hat{\mathbf{g}}_i^{[j+1]} = \frac{1}{(\mathbf{B}^{-1})_{ii}} \frac{1}{k} \mathbf{Z}' \mathbf{A}_{22}^{-1} \hat{\mathbf{a}}_{2(-i)}^{[j]} = \frac{1}{(\mathbf{B}^{-1})_{ii}} \frac{1}{k} \mathbf{Z}' \hat{\boldsymbol{\gamma}}^{[j]}$$

- Step 2. update residual polygenic effects

$$\hat{\mathbf{a}}_{2(+i)}^{[j]} = \hat{\mathbf{a}}_{2(-i)}^{[j]} + \mathbf{z}_i (\hat{\mathbf{g}}_i^{[j]} - \hat{\mathbf{g}}_i^{[j+1]})$$

